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RESEARCH

The impact of antibiotics on growth in children in low and middle income countries: systematic review and meta-analysis of randomised controlled trials

 OPEN ACCESS

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Abstract

Objectives To determine whether antibiotic treatment leads to improvements in growth in prepubertal children in low and middle income countries, to determine the magnitude of improvements in growth, and to identify moderators of this treatment effect.

Design Systematic review and meta-analysis.

Data sources Medline, Embase, Scopus, the Cochrane central register of controlled trials, and Web of Science.

Study selection Randomised controlled trials conducted in low or middle income countries in which an orally administered antibacterial agent was allocated by randomisation or minimisation and growth was measured as an outcome. Participants aged 1 month to 12 years were included. Control was placebo or non-antimicrobial intervention.

Results Data were pooled from 10 randomised controlled trials representing 4316 children, across a variety of antibiotics, indications for treatment, treatment regimens, and countries. In random effects models, antibiotic use increased height by 0.04 cm/month (95% confidence interval 0.00 to 0.07) and weight by 23.8 g/month (95%

confidence interval 4.3 to 43.3). After adjusting for age, effects on height were larger in younger populations and effects on weight were larger in African studies compared with other regions.

Conclusion Antibiotics have a growth promoting effect in prepubertal children in low and middle income countries. This effect was more pronounced for ponderal than for linear growth. The antibiotic growth promoting effect may be mediated by treatment of clinical or subclinical infections or possibly by modulation of the intestinal microbiota. Better definition of the mechanisms underlying this effect will be important to inform optimal and safe approaches to achieving healthy growth in vulnerable populations.

Introduction

Undernutrition in early childhood, characterised by poor linear or ponderal growth, underlies approximately one third of all mortality in children aged under 5 years worldwide.¹ Linear growth, measured as height or length, is an indicator of long term nutritional status; children whose height for age is more than 2 standard deviations below the reference population mean

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Search strings
Review protocol

are termed stunted. Ponderal growth, measured as body weight, is viewed as an indicator of short term or long term nutritional status. Children whose weight for age is more than 2 standard deviations below the reference population mean are termed underweight. Underweight and stunting, particularly during the first two years of life, have short term effects on morbidity and mortality and long term effects on cognition, educational achievement, and economic productivity as an adult.² Given the current global focus on reducing underweight and stunting to reach forthcoming global health targets,^{3,4} interest in evaluating interventions to promote healthy growth in childhood is increasing.⁵ Primary interventions to improve growth in children have largely focused on nutritional supplementation and prevention of diarrhoea. However, the impact of these interventions on restoring growth deficits in undernourished children is modest.⁶⁻⁹ Restoration of deficits in linear growth is particularly challenging beyond the first two years of life.²

The growth promoting effects of antibiotics were first observed in animals in the 1940s. Small daily doses of broad spectrum antibiotics have been found to improve average daily weight gain in farm animals by as much as 73%.¹⁰⁻¹⁸ These observations led to the hypothesis that food animals reared in conditions of poor sanitation and hygiene have impaired growth because of chronic exposure to environmental microbes and pathogens, and that antibiotic treatment may therefore improve growth.¹⁹

In humans, an association between infections and malnutrition in children is supported in the literature.^{20,21} Nutrient harvesting from the diet and the inflammatory response of the gut are also modulated by the intestinal microbiota, a microbial ecosystem that is essential to human health and nutrition.²²⁻²⁶ Perturbation of this microbial community through chronic exposure to environmental microbes or pathogens may also be detrimental to growth in children,^{19,27-29} and studies have shown that antibiotic use can affect the composition of the microbial community.^{30,31} Antibiotic use has also been associated with significant height and weight gains among children in some target populations.³²⁻³⁵ However, results have not always been consistent,^{32,36-39} and researchers continue to investigate the potential co-benefits of antibiotic treatment on growth in children.^{40,41}

We carried out a systematic review of randomised controlled trials to determine whether improvements in growth are seen among prepubertal children (1 month to 12 years old) treated with antibiotics in low and middle income countries; to determine the magnitude of these growth effects; and to identify moderators of this treatment effect. We hypothesised that antibiotics would have a positive average effect on both height and weight, and that treatment effect size would be moderated by the characteristics of antibiotic treatment, differences in study population, and trial quality.

Methods

Search strategy and selection criteria

This review is reported in accordance with the PRISMA statement⁴² and recommendations for reporting meta-analyses of individual patient data.⁴³ We searched Medline (including In-Process and Other Non-Indexed Citations) and Embase, both using Ovid, as well as Scopus and the Cochrane central register of controlled trials up to 12 December 2013. A professional librarian helped to develop the search strings (see supplementary appendix 1 for details of search strings and appendix 2 for the review protocol).

We searched for randomised controlled trials conducted in low or middle income countries with participants aged 1 month to 12 years allocated by randomisation or minimisation to antibacterial treatment given by mouth, or to control. Control interventions included placebo, an intervention with no known antimicrobial effect, or no treatment. We selected trials, published or unpublished, if growth was measured as an outcome. Studies of anthelmintic treatments were excluded, since systematic reviews of such trials have already been conducted.^{44,45} We placed no restrictions on language, year of publication, or the length of follow-up, and excluded quasiexperimental studies, observational studies, reviews, and simulations. We excluded studies of neonates (<1 month old) since growth patterns during the neonatal period, particularly among preterm infants, are different from the post-neonatal period. Finally, we considered trials ineligible for inclusion if the condition being treated did not depend on the antimicrobial effect of antibiotic treatment (for example, use of specific antibiotics to reduce feeding intolerance through prokinetic effects or to improve lung function through anti-inflammatory effects).

Two investigators (EKG and SMAJ) independently assessed titles and abstracts for eligible publications. If eligibility could not be determined, the full article was retrieved and the methods screened. In an effort to find similar trials we used Web of Science to search for publications that cited the included studies, and we also handsearched reference lists of included trials and any review articles identified. A third investigator (AJP) adjudicated discrepancies.

Data abstraction and analysis

Study quality was determined by assessing the included publications for risk of bias from the procedures used for sequence generation, allocation concealment, and blinding; and by informative censoring or selective outcome reporting using a standardised instrument adapted from the Cochrane handbook.⁴⁶ Two reviewers (EKG and SMAJ) independently assessed the included publications. Discrepancies were resolved by consensus.

We contacted study authors up to three times by email (or by telephone if email was unsuccessful) to determine their interest in collaborating on this review and to request individual patient data. When such data could not be obtained, the same two reviewers independently abstracted data using a standardised pretested form, with discrepancies resolved by consensus. For each trial arm we abstracted number of participants, number lost to follow-up or excluded after randomisation, mean baseline height or weight, and mean height or weight (and standard deviations) at the end of follow-up. For reported treatment effects we also abstracted P values, confidence intervals, and standard errors. Where mean change in height or weight for each unit of follow-up time was reported, we retrieved the same information. We also abstracted several trial level characteristics, which we defined a priori as potential moderators of treatment effect: indication for treatment, country, proportion boys, mean age, antibiotic agent, dosage, frequency and duration of antibiotic treatment, concurrent interventions, length of follow-up, and whether treatment effects were adjusted for imbalances at baseline. We defined antibiotic class as bacteriostatic or bactericidal, and antibiotic spectrum as broad or narrow. Broad spectrum antibiotics were defined as those reported in the literature to be effective against a wide range of Gram positive and Gram negative bacteria, and narrow spectrum antibiotics as those reported in the literature to be effective

against a limited range of bacteria.⁴⁷⁻⁵⁵ Risk of bias domains were treated as potential sources of heterogeneity.

Outcomes were mean height (cm) or weight (g) at the end of follow-up or mean change in height (cm) or weight (g) per unit of follow-up time. The difference in means between treatment and control arms was the measure of treatment effect. We scaled the treatment effects and their variances as the average effect for each month of follow-up. We analysed height and weight separately. When a trial allocated participants to more than one intervention, we abstracted data from the arm allocated to receive antibiotics.⁵⁶⁻⁵⁸ When a trial allocated participants to more than one antibiotic arm, we combined the data from both arms to avoid unit of analysis errors.^{34 59}

We combined individual patient data and aggregate data using a two step approach.⁶⁰ In the first step we estimated treatment effects in each trial with individual patient data in an intention to treat analysis using linear mixed models to allow for random intercepts and serial correlation. For each trial we fit one individual patient data model, with baseline growth, age, sex, duration of follow-up, and a duration by treatment interaction included as covariates. In the second step we used a random effects model to pool intention to treat effect estimates (obtained from separate individual patient data trials in step 1) with intention to treat effect estimates abstracted from publications with aggregate data.

We assessed statistical heterogeneity using the I^2 statistic.⁶¹ Heterogeneity was explored using weighted metaregression and subgroup meta-analyses. Statistical significance was evaluated at $\alpha < 0.05$. We assessed publication bias using Egger's test.⁶¹ Sensitivity analyses were performed in two ways: to determine the robustness of meta-analysis results to the removal of studies,⁶² and by fitting linear mixed models restricted to the five trials for which individual patient data were available.

Trials with individual patient data were modelled using the lme4 package, and we fit all meta-analyses and metaregression models using the metafor package,⁶³ both using R version 2.15.1.

Results

Study selection

The electronic search identified 4600 records. An additional 24 records were identified through Web of Science and a backward search of reference lists (fig 1). Of these, 190 studies were retrieved and screened for eligibility. Overall, 139 studies failed to meet at least two selection criteria. Thirty four additional studies were excluded for failing to meet one of the following criteria: no antibiotic was allocated ($n=8$), an active comparator was used ($n=7$), the trial was not conducted in a low or middle income country ($n=5$), treatment was not randomised ($n=5$), growth was not measured or reported ($n=4$), participants' age range exceeded 12 years ($n=2$), review articles ($n=2$), and the antibiotic was not administered orally ($n=1$). Only four were non-English language texts. These were screened using an electronic translator.

Four additional trials were excluded because they reported differences in prevalence of stunting or wasting,⁶⁴⁻⁶⁶ or reported growth using the Wetzel grid method.⁶⁷ Of these, three authors could not be reached to request individual patient data or unpublished data,^{64 66 67} and data were no longer available from one.⁶⁵ Another author was contacted and provided individual patient data but did not provide a data dictionary. Since this publication did not report outcomes for the antibiotic arms, this trial was excluded.⁶⁸ Of these five otherwise eligible randomised controlled trials that were excluded because growth was not

reported in the desired format, four reported no growth benefits from antibiotics,^{64-66 68} but they would only have represented 8% of the total person time if they were included and would not have greatly influenced our findings. Another two trials were excluded because they only reported growth at baseline and the authors could not be reached.^{69 70}

Published data were available from five trials,^{56 58 59 71 72} and complete datasets for individual patient data were obtained from five trials.^{33-35 57 73} Thus 10 randomised controlled trials were included in the meta-analysis.^{33-35 56-59 71-73} Only data from the secnidazole and placebo arms were included for Goto and colleagues,⁵⁷ from the metronidazole and placebo arms for Gupta and colleagues,⁵⁸ and from the metronidazole and no intervention arms for Heikens and colleagues.⁵⁶

Study characteristics

Of these 10 trials, nine were placebo controlled and one gave the controls no treatment.⁵⁶ Indication for treatment varied by trial and included malnutrition ($n=4$), infection with *Giardia lamblia* ($n=2$), diarrhoea with or without vomiting ($n=2$), environmental enteropathy ($n=1$), and prophylaxis in children infected with human immunodeficiency virus ($n=1$). The earliest included trial was published in 1953 and the most recent in 2013. Three trials gave a nutritional supplement to participants in both arms^{34 56 71} in addition to antibiotics or control (table 1). Only three trials with aggregate data reported the number of male participants (table 2). Two trials recruited children admitted to hospital, and both reported weight only.^{33 72}

Eight trials reported height^{34 35 56-59 71 73} and all reported weight.^{33-35 56-59 71-73} Four trials with individual patient data reported height^{34 35 57 73} and five reported weight.^{33-35 57 73} Together these trials included 1699 control and 2617 antibiotic treated participants, followed-up for a mean of 268 (SD 266) days, across seven countries. The mean age of participants ranged from 4 to 115 months (table 2). On average, trial participants were below the age standardised reference population mean for height or weight at baseline (table 1).

Risk of bias

Only one trial⁵⁹ was evaluated to be at high risk for bias overall (that is, when all bias domains were considered together). This was based on high risk as a result of inadequate random sequence generation, since treatment was randomly allocated to groups of children determined by the investigators; unclear risk due to inadequate allocation concealment; and high risk from differential attrition between treatment arms. Five trials were ranked as low risk for bias overall. These trials were low risk in all six bias domains.^{33-35 57 73} Finally, four trials had an unclear risk for bias overall because the procedures were not fully described.^{56 58 71 72}

For risk of bias due to attrition, the linear mixed models fit for the five trials with individual patient data were unbiased by losses to follow-up, provided the losses were uninformative conditional on observed height and weight. Among the trials with aggregate data only Guzman and colleagues⁵⁹ was determined as being impacted by drop outs. Gupta and colleagues⁵⁸ reported exclusions before randomisation but reported outcomes on all 79 participants recruited at baseline. Heikens and colleagues⁵⁶ reported that drop outs predominantly consisted of participants who moved too far from the study site to be followed-up or withdrew consent (9% and 2% of the total sample, respectively). Risk of bias due to attrition could not be assessed in the two remaining trials with aggregate data because the authors did not provide any data on exclusion of participants

during the study.^{71 72} However, these trials only accounted for 5.9%⁷¹ of the weights in the pooled treatment effect for height, and 1.2%⁷¹ and 1.4%⁷² in the pooled treatment effect for weight. Overall, we do not think that attrition posed a serious risk of bias in our analyses.

Egger's test suggested no significant publication bias among trials reporting height ($P=0.841$) or weight ($P=0.391$).

Meta-analysis

Our random effects models estimated an average treatment effect for height of 0.04 cm/month (95% confidence interval 0.00 to 0.07, fig 2), and an average effect for weight of 23.8 g/month (95% confidence interval 4.3 to 43.3) in antibiotic treated compared with control children (fig 3). The I^2 statistic showed a considerable degree of statistical heterogeneity in treatment effects for both height and weight (84.8% and 84.4%, respectively).

To assess the impact of antibiotic treatment on growth in children aged less than 2 years, we used the same two step approach described for the analysis using complete data. We fitted models of individual patient data restricted to participants less than 2 years old^{33-35 57} and pooled these treatment effect estimates with the estimates for aggregated data reported by Heikens and colleagues,⁵⁶ which was the only trial with aggregated data restricted to this age group. These included observations from 833 control and 1461 treated infants, followed up for a mean 169 (SD 152) days. The treatment effect in these children was not statistically significant for height (0.03 cm/month, 95% confidence interval -0.05 to 0.11) but was for weight (29.6 g/month, 95% confidence interval 2.4 to 56.8), $I^2=47.0\%$.

Metaregression analyses

Only geographical region significantly explained variation in the treatment effect across trials for weight (table 3). The treatment effect was 35.6 g/month larger on average in trials conducted in Africa (95% confidence interval 12.8 to 58.3) compared with trials conducted in other regions. No statistically significant moderators of the height treatment effect were identified by bivariate analyses. We could not investigate risk of bias domains as moderators of treatment effect because only one trial was evaluated as high risk in any domain. All bivariate models included one treatment effect moderator and one outcome (table 3).

Duration of treatment, geographical region, treatment for *Giardia lamblia* infection, and age were statistically significant moderators of treatment effect, after adjustment for mean age of study population (table 4). The height treatment effect was 0.001 cm/month (95% confidence interval -0.002 to 0.000) smaller on average with each one month increase in mean population age, and was 0.007 cm/month larger on average with each additional day of treatment (0.00 to 0.01). The weight treatment effect was 0.5 g/month smaller on average (95% confidence interval -1.0 to -0.1) with each one month increase in mean age, 33.2 g/month (5.3 to 61.2) larger on average in trials conducted in Africa, and 46.9 g/month (-83.2 to -10.6) smaller on average in trials in which participants were treated for *G lamblia* infection. In this last model, the intercept was 62.1 g/month (95% confidence interval 29.3 to 94.9), indicating a significant treatment effect in trials that did not treat children for *G lamblia* infection. All mean age adjusted models included mean participants' age, one treatment effect moderator, and one outcome (table 4).

Sensitivity analyses

Only removal of Prendergast and colleagues³⁵ from the random effects model impacted the average effect for height. Without this trial the average effect was 0.02 cm/month (95% confidence interval -0.01 to 0.05), a 50% decrease. The average treatment effect for weight was robust to the removal of trials. Also, two trials recruited children admitted to hospital.^{33 72} Simultaneous exclusion of both trials did not change the average treatment effect for weight (21.5 g/month, 95% confidence interval 2.3 to 40.7). These two trials did not report height.

In addition we fit linear mixed models to investigate whether adjusting for participant age at the individual level (using trials with individual patient data only) would produce the same estimates of treatment effect moderation as we obtained by weighted metaregression adjusted for mean participant age (table 4). These models included age and duration of treatment, geographical region, or treatment for *G lamblia* infection, along with corresponding interaction terms. Results of these individual patient data models were consistent with the weighted metaregression results using all trials, with the exception of age, where the treatment effect on weight increased by 0.8 g/month for each one month increase in child age on average.

Subgroup analyses

The weight treatment effect was homogeneous across trials conducted in Africa using a random effects model (41.4 g/month, 95% confidence interval 31.0 to 51.7); $I^2=0.0\%$. The average treatment effect estimated in this subgroup was identical when a fixed effects model was used (41.4 g/month, 31.0 to 51.7).

Discussion

In this pooled analysis of individual patient data and aggregate data from 10 randomised controlled trials conducted in seven low and middle income countries, antibiotic treatment had a positive average treatment effect on both height and weight in children aged 1 month to 12 years. Our results suggest that the growth promoting effect of antibiotics is more substantial for ponderal growth than for linear growth, and that the effect may be more homogenous in younger children. Analysis from the trials with individual patient data showed an increase in the weight treatment effect with increasing participants' age. This is in contrast with the results of the metaregression model, which suggested a smaller effect with increasing mean age. The trials with individual patient data primarily included children less than 5 years old, whereas two trials with aggregate data recruited older children.^{51 79} Cross level bias may also partly explain this discrepancy. Although we did not restrict study selection to populations with a particular nutritional status, children were generally below the age standardised reference population mean for height or weight, reflecting the spectrum of stunting and wasting malnutrition seen in low and middle income countries. The larger weight treatment effects we observed in trials conducted in Africa may plausibly be explained by the high prevalence of HIV infection and severe acute malnutrition among populations included in these studies. Two trials conducted in Africa included severely malnourished children in whom all or a subset were infected with or exposed to HIV.^{34 35} A third trial also included children from a similar high HIV prevalence community,^{34 73} although HIV status was not specifically reported. The smaller weight treatment effect we observed in trials treating children for *G lamblia* infection suggested that growth may not be as strongly impacted in children treated with antibiotics for this specific protozoal

infection. Overall, the average treatment effects we observed would correspond to an approximate 0.1 increase in height for age Z score and a 0.2 to 0.3 increase in weight for age Z score over six months in HIV infected, HIV exposed, or severely malnourished populations under 2 years old using the World Health Organization growth standard.⁷⁴ These treatment effects therefore represent clinically relevant growth gains for the youngest children from the most vulnerable populations, in whom the long term impact of undernutrition is most profound.

Strengths and limitations of this study

The inclusion of trials with individual patient data and aggregate data served to improve the precision of our pooled estimates, minimised the risk of publication bias,⁶⁰ and allowed us to define height and weight in uniform units, avoiding the potential sources of bias inherent in utilising standardised mean differences.⁷⁵ We synthesised data from 4316 children, observed across a variety of antibiotics, indications for treatment, treatment regimens, and countries, providing the first comprehensive review of evidence from all randomised trials relating antibiotic use to growth in children in low and middle income countries, conducted over a 60 year period. A clear limitation of pooling such a diverse set of trials, with a large degree of statistical heterogeneity, is the limited generalisability of the average treatment effects. It is not completely clear which antibiotics or treatment regimens can be expected to produce these growth effects in other populations. However, pooling this diverse set of trials did allow identification of important subpopulations in whom the growth effect may be more profound when broad spectrum antibiotics are used. However, owing to the small number of trials, we had limited power to identify moderators of treatment effect, and we were not able to fully investigate trial level confounding with multivariable meta-regression models. Specifically, the potential modifying effect of HIV prevalence, treatment duration, antibiotic class, concurrent nutritional interventions, and study population characteristics could not be fully elucidated. Also, cross level bias cannot be ruled out in our meta-regression analyses of treatment effect moderators (which are measured at trial level); hence care must be taken in extending the treatment modifying effects to the individual level, particularly for age. Egger's test showed no evidence of publication bias. Careful screening of search results and communication with investigators ensured identification of published and unpublished reports. Finally, only one trial was evaluated to be at high risk for bias.⁵⁹

Comparison with other studies

The exact reasons for the observed growth effects from antibiotics remain unclear, but several mechanisms may be involved. Respiratory and gastrointestinal infections are known to be associated with undernutrition. Malabsorption of nutrients, increased nutrient loss during episodes of diarrhoea, gut inflammation, impaired intestinal barrier function, diversion of nutrients away from growth to support immune activation, and loss of appetite are possible reasons for impaired growth during infection.¹⁹⁻²¹ Antibiotics may improve growth by resolving subclinical and clinical infections. Eradication of microbes that regulate endocrine hunger signals may also contribute to growth gains with antibiotics. Changes in post-meal leptin and ghrelin serum levels, both of which help to regulate satiety, have been associated with the eradication of *Helicobacter pylori* following antibiotic treatment,⁷⁶ although this may play less of a role in food insecure settings.

An alternative possibility is that alteration of the intestinal microbiota by antibiotics may result in growth gains.⁷⁷⁻⁷⁹ The

intestinal microbiota regulates immune development and inflammation in the gut,²³⁻²⁴ maintains host-microbe homeostasis in the gut,²⁵ and has an important role in nutrient harvesting and absorption.²⁶ Disturbance of intestinal microbiota composition resulting from chronic intestinal colonisation with pathogens or overgrowth of commensal bacteria in the small intestine¹⁹⁻²⁷⁻²⁹ may lead to disruption of these functions. Perturbation of the intestinal microbiota may also lead to intestinal inflammation and increased intestinal permeability. These changes are characteristic of environmental enteropathy, a subclinical disorder of the intestinal tract that is ubiquitous in developing countries and is associated with poor linear growth.¹⁹⁻²⁷⁻²⁹

Antibiotics are known to induce changes in the composition of microbiota in the gut,³⁰⁻³¹ and these changes may persist.⁷⁷⁻⁷⁹ Recent work has shown that intestinal microbial taxa may not return to their pretreatment abundance levels, even after a single use of antibiotics⁷⁷⁻⁷⁹⁻⁸⁰; however, the extent of recovery to baseline may depend on the class of antibiotic used.⁷⁹ A recent review qualitatively summarised the evidence supporting a relation between antibiotic use and weight and included evidence from some trials in humans.⁷⁷ The mechanisms underlying these growth benefits plausibly include resolution of underlying infections or inflammatory processes (for example, environmental enteropathy) or alteration of intestinal microbiota composition and function. In an experimental animal model, weight loss in mice resulted from transplantation of donor faeces from children with kwashiorkor, but not from their healthy twins,⁸¹ whereas increases in total body mass and fat mass were induced in mice transplanted with donor faeces from obese adults, but not from their lean twins.⁸² Although we cannot rule out an effect of antibiotics on latent bacterial infections in the included trials, it is plausible that the growth benefits we observed also encompass an important growth effect mediated by intestinal microbiota.

Conclusions and policy implications

In summary, our results show that antibiotic treatment has a growth promoting effect, particularly for ponderal growth, in prepubertal children from undernourished populations in low and middle income countries. Linear growth seems less responsive to antibiotics. A better understanding of the biological mechanisms behind these antibiotic associated effects on growth is critical for certain populations, such as children under 2 years old (as reversal of stunting beyond this age is challenging²), and HIV infected, HIV exposed, and acutely malnourished children in whom antibiotics continue to be a standard component of care.³⁶⁻⁸³⁻⁸⁴ Antibiotics, however, are not the most viable option for the treatment of malnutrition outside of these highly vulnerable populations in which antibiotic treatment is already routinely recommended for treatment and prevention of infections. In addition to concerns about antimicrobial resistance, antibiotic use has also been associated with adverse events such as antibiotic associated diarrhoea. The growth benefits of more widespread antibiotic use may not outweigh the risks. Our findings highlight the co-benefits of antimicrobial treatment that have been previously reported from developing countries³⁴⁻³⁵ and provide an intriguing proof of concept that treatment of subclinical infections and modulation of the intestinal microbiota may have beneficial effects on growth.

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What is already known on this topic

Antibiotics have shown variable effects on growth in humans but are currently recommended for severely malnourished children and those infected with or exposed to HIV to reduce morbidity and mortality

Several mechanisms exist through which antibiotic treatment may affect growth in children, including resolution of infection and, potentially, alteration of the intestinal microbiota

What this study adds

Evidence from a diverse set of randomised controlled trials show that antibiotic use in prepubertal children from undernourished populations in low and middle income countries leads to clinically relevant growth gains, particularly for weight

Larger growth gains are associated with antibiotic use in studies with a high prevalence of HIV infection and severe acute malnutrition

The growth gains show the co-benefits of antibiotic treatment in high risk populations, and provide proof of concept that treatment of infections or modulation of the intestinal microbiota can have beneficial growth effects; however, more research is needed to better understand the mechanisms involved

Contributors: EKG, EEEM, JHH, RJS, and ARM created and designed the study. EKG and SMAJ conducted the literature search, study selection, and data collection. AJP adjudicated study selection. AJP, DMG, ASW, IT, RG, ST, and MBdM provided individual patient data, and provided insight into trial design, implementation, and database structure. EKG did the data analysis. EKG, EEEM, and ARM interpreted the data and drafted the manuscript. All authors critically revised the manuscript for intellectual content, discussion of findings, and overall conclusions. ARM is the guarantor.

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Ethical approval: Not required.

Transparency: ARM affirms that this manuscript is an honest, accurate, and transparent account of the study being reported; that no important aspects of the study have been omitted; and that any discrepancies from the study as planned have been explained.

Data sharing: The electronic data abstraction form is available from the first author at ethan.gough@mail.mcgill.ca.

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Tables

Table 1 | Characteristics of randomised controlled trials of antibiotic use and growth in prepubertal children included in meta-analysis

| Study, country | Indication for treatment | Eligibility criteria | Baseline nutritional status | Intervention | | |
|---|---|---|---|---------------------------|---------|--|
| | | | | Antibiotic | Control | Concurrent |
| Scrimshaw et al 1953 ⁷¹ , Guatemala | Malnutrition | Schoolchildren | Children in participating communities averaged 2-4 years below US reference for height and weight | Aureomycin | Placebo | Enriched soya milk powder given 6 days/week except during holidays |
| Guzman et al 1958 ⁵⁹ , Guatemala | Malnutrition | Schoolchildren | Children in participating communities averaged 2-4 years below US reference for height and weight | Aureomycin or penicillin | Placebo | None |
| Wolfsdorf et al 1973 ⁷² , South Africa | Diarrhoea with or without vomiting | Infants presenting with diarrhoea or vomiting severe enough to warrant hospital stay | Not recorded | Trimethoprim-sulphonamide | Placebo | "Routine" treatment regimens carried out |
| Gupta et al 1982 ⁵⁸ , Guatemala | <i>Giardia lamblia</i> | Children | Mean percentage height and weight for age: 88.6% and 71.5% | Metronidazole | Placebo | None |
| Heikens et al 1993 ⁵⁶ , Jamaica | Malnutrition | Children malnourished according to Wellcome classification, excluding children with oedema, congenital abnormality, infection requiring hospital stay, or anorexia preventing normal home feeding | Mean percentage height and weight for age: 88.6% and 65.1% | Metronidazole | None | Multivitamins and folic acid, outpatient treatment of infection or illness, advice on breast feeding and weaning for duration of follow-up |
| Tahan et al 2007 ⁷³ , Brazil | Diarrhoea | Infants with diarrhoea for at least 7 days who needed hospital stay, excluding infants with associated disorders, use of antibiotics in preceding 7 days, or evidence of systemic infection | Mean height and weight for age Z scores: -2.02 and -2.36 | Polymixin B | Placebo | None |
| Goto et al 2009 ⁵⁷ , Bangladesh | <i>G lamblia</i> | Infants | Mean height and weight for age Z scores: -1.05 and -1.82 | Secnidazole | Placebo | None |
| Trehan et al 2009 ⁷⁴ , Malawi | Environmental enteropathy | Children, excluding those with chronic debilitating illnesses or evidence of severe acute malnutrition | Mean height and weight for age Z scores: -1.67 and -0.91 | Rifaximin | Placebo | None |
| Prendergast et al 2011 ³⁵ , Zambia | Prophylaxis against opportunistic infection | Children with positive HIV antibody test result, excluding those with opportunistic infection, life expectancy \leq 4 weeks, current cotrimoxazole treatment or allergy to this drug, or previous <i>Pneumocystis jirovecii</i> pneumonia | Mean height and weight for age Z scores: -3.55 and -3.10 | Cotrimoxazole | Placebo | None |
| Trehan et al 2013 ³⁴ , Malawi | Severe acute malnutrition | Children with oedema, or weight for height Z score \leq 3 | Mean height for age Z score was -3.19 | Amoxicillin or cefdinir | Placebo | Standardised nutrition counselling and ready to use therapeutic food at dose of approximately 175 kcal/kg/day given in 2 week intervals |

Table 2| Growth outcomes and potential treatment effect moderators in randomised controlled trials of antibiotic use and growth in prepubertal children included in meta-analysis

| Study | IPD | Mean (SD) age (months) | No (%) male | Antibiotics | | | | Mean follow-up (days) | Mean growth/month of follow-up | | | |
|---------------------------------|-----|------------------------|-------------|--------------------------------|---|-----------|--------------|-----------------------|--------------------------------|---------|------------|-----------|
| | | | | Class, spectrum | Dosage | Doses/day | Days treated | | Height (cm) | | Weight (g) | |
| | | | | | | | | | Controls | Treated | Controls | Treated |
| Scrimshaw et al ⁷¹ | No | 114.9 (NR*) | 143 (57.2)* | Bacteriostatic, broad spectrum | 50 mg | 1 | 667† | 758 | 0.39 | 0.42 | 180.0 | 270.0 |
| Guzman et al ⁵⁹ | No | 114.9 (NR) | 143 (57.2) | Bacteriostatic, broad spectrum | 50 mg | 1 | 394† | 394 | 0.36 | 0.36 | 170.0 | 166.0 |
| Wolfsdorf et al ⁷² | No | 5.9 (6.4) | NR | Bactericidal, broad spectrum | NR | NR | NR | 91 | NR | NR | 664.0 | 788.4 |
| Gupta et al ⁵⁸ | No | 23.0 (17.2‡) | NR | Bactericidal, narrow spectrum | 25 mg/kg | 2 | 42 | NR | 0.51 | 0.58 | 135.9 | 154.2 |
| Heikens et al ⁵⁶ | No | 14.1 (6.5) | NR | Bactericidal, narrow spectrum | 20 mg/kg | 1 | 5 | 179 | 12.40 | 12.20 | 1336.7 | 1393.3 |
| Tahan et al ³³ | Yes | 4.0 (2.0) | 17 (68.0) | Bactericidal, narrow spectrum | 2.5 mg/kg | 4 | 7 | 7 | NR | NR | 710.5§ | 735.7§ |
| Goto et al ⁵⁷ | Yes | 8.6 (3.2) | 135 (50.4) | Bactericidal, narrow spectrum | 35 mg/kg | 1 | 10 | 264 | 9.11 | 9.12 | 1105.1 | 1100.7 |
| Trehan et al ⁷³ | Yes | 47.2 (7.12) | 60 (41.7) | Bacteriostatic, broad spectrum | 10 mg | 2 | 7 | 28 | 107.56¶ | 107.96¶ | 14 957.1¶ | 15 030.3¶ |
| Prendergast et al ³⁵ | Yes | 64.5 (44.7) | 266 (49.2) | Bactericidal, broad spectrum | 240 g (<5 yrs); 480 g (>5 yrs) | 1 | 575 | 575 | 5.66 | 5.77 | 803.9 | 845.0 |
| Trehan et al ³⁴ | Yes | 21.1 (9.1) | 1317 (47.6) | Bactericidal, broad spectrum | 7 mg/kg (cefdinir); 40-45 mg/kg (amoxicillin) | 2 | 7 | 43 | 26.74 | 26.76 | 2898.0 | 2938.9 |

IPD=individual patient data; NR=not reported.

*Not reported by Schrimshaw et al.⁷¹ Values assumed to be same as in Guzman et al⁵⁹ as both studies were conducted in communities in Guatemalan highlands in 1950s by same research group and recruited children in 5-12 year age range.

†Estimated from mean number of treatment days reported per trial arm.

‡Not reported by Gupta et al,⁵⁸ estimated from Schrimshaw et al 1968.⁸⁵

§Mean change in weight per day; follow-up was seven days.

¶Follow-up was 28 days; these represent height and weight at end of follow-up.

Table 3| Estimated average differences in antibiotic treatment effects on growth in prepubertal children, using weighted bivariate random effects metaregression

| Trial characteristics | No | Height (cm/month) | | | No | Weight (g/month) | | |
|---|----|-------------------|---------|-----------------------------|----|------------------|---------|-----------------------------|
| | | Mean difference | P value | I ² (%) (95% CI) | | Mean difference | P value | I ² (%) (95% CI) |
| Geographical region (Africa v other) | 8 | 0.05 | 0.275 | 79.6 (39.7 to 98.8) | 10 | 35.57 | 0.002 | 50.9 (10.6 to 99.1) |
| Publication year | 8 | 0.00 | 0.650 | 78.7 (35.5 to 99.0) | 10 | 0.50 | 0.275 | 77.6 (47.1 to 99.9) |
| Treatment effect adjusted for baseline imbalances (yes v no) | 8 | 0.00 | 0.964 | 88.5 (55.9 to 99.9) | 10 | -17.59 | 0.465 | 85.4 (62.8 to 99.9) |
| Mean length of follow-up (days) | 7 | 0.00 | 0.282 | 79.8 (35.3 to 99.4) | 9 | -0.05 | 0.490 | 87.5 (61.1 to 99.9) |
| No of doses/day | 8 | 0.02 | 0.648 | 83.4 (44.1 to 98.8) | 9 | 21.70 | 0.307 | 84.4 (56.9 to 99.7) |
| Duration of treatment (days) | 8 | 0.00 | 0.340 | 81.9 (47.1 to 99.1) | 9 | 0.00 | 0.921 | 86.6 (65.6 to 99.9) |
| Antibiotic class (bactericidal v bacteriostatic)* | 7 | -0.05 | 0.792 | 71.6 (25.9 to 99.6) | 8 | -51.61 | 0.727 | 87.1 (72.1 to 100.0) |
| Antibiotic spectrum (broad v narrow) | 8 | 0.02 | 0.666 | 89.2 (57.2 to 99.3) | 10 | 9.41 | 0.666 | 84.6 (61.0 to 99.9) |
| Participants given concurrent nutritional intervention (yes v no) | 8 | -0.05 | 0.356 | 82.5 (44.9 to 98.6) | 10 | 31.00 | 0.110 | 75.7 (44.2 to 99.9) |
| Mean age (months) | 8 | 0.00 | 0.948 | 82.0 (1.7 to 96.7) | 10 | -0.24 | 0.381 | 82.0 (54.3 to 99.9) |
| Treatment was for malnutrition (yes v no) | 8 | -0.06 | 0.066 | 75.2 (17.4 to 99.0) | 10 | 2.65 | 0.906 | 85.1 (62.5 to 99.9) |
| Treatment was for <i>Giardia lamblia</i> infection (yes v no) | 8 | 0.01 | 0.833 | 88.6 (56.2 to 99.4) | 10 | -26.42 | 0.210 | 82.2 (55.1 to 99.9) |
| Treatment was for diarrhoea with or without vomiting (yes v no)† | NA | NA | NA | NA | 10 | 144.37 | 0.075 | 85.3 (60.9 to 99.8) |

NA=not applicable.

*Excludes Prendergast et al³⁵ and Wolfsdorf et al⁷² as not clear whether trimethoprim with sulphonamide or sulfamethoxazole are bacteriostatic or bactericidal in combination.

†No trials reporting height treated participants for diarrhoea with or without vomiting.

Table 4| Significant moderators of antibiotic treatment effects on growth in prepubertal children, using weighted random effects metaregression adjusted for mean study population age

| Trial characteristics | No | Mean difference (95% CI) | I ² (%) (95% CI) |
|---|----|--------------------------|-----------------------------|
| Height model 1 (cm/month): | | | |
| Duration of treatment (days) | 8 | 0.007 (0.00 to 0.01) | 53.6 (0.0 to 99.3) |
| Mean age (months) | 8 | −0.001 (−0.002 to 0.00) | |
| Weight model 1 (g/month): | | | |
| Geographical region (Africa v other) | 10 | 33.2 (5.3 to 61.2) | 53.5 (3.6 to 99.9) |
| Mean age (months) | 10 | −0.2 (−0.4 to −0.1) | |
| Weight model 2 (g/month) | | | |
| Treatment was for <i>Giardia lamblia</i> (yes v no) | 10 | −46.9 (−83.2 to −10.6) | 57.8 (9.3 to 99.9) |
| Mean age (months) | 10 | −0.5 (−1.0 to −0.1) | |

Figures

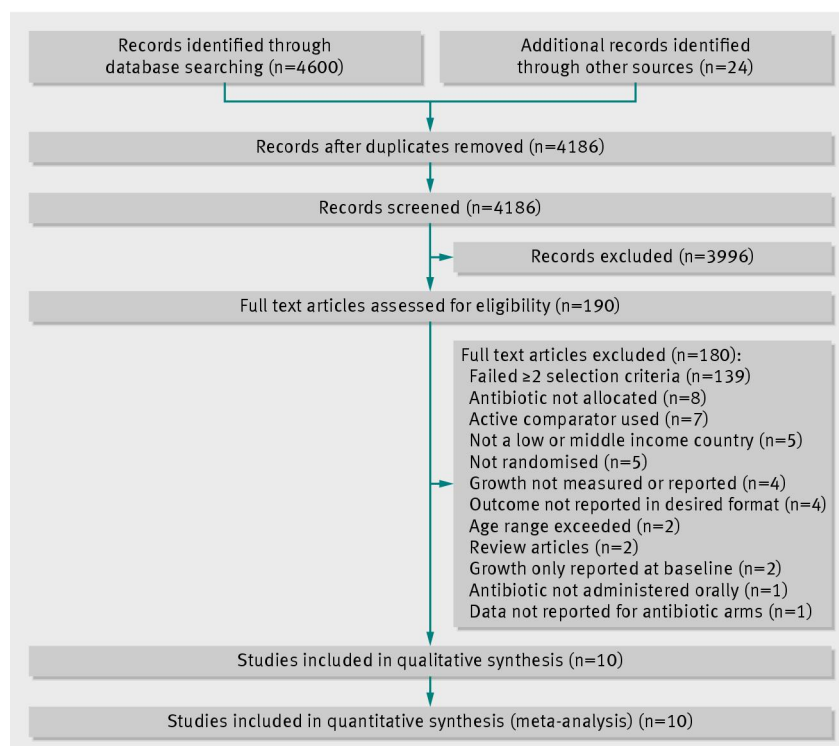


Fig 1 Flow diagram of search retrieval and trial selection

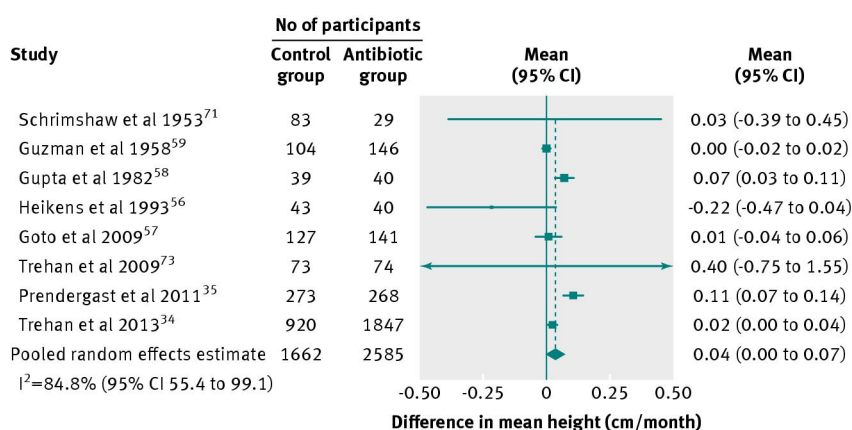


Fig 2 Random effects meta-analyses and forest plots of antibiotic use and height. Point size reflects study weight

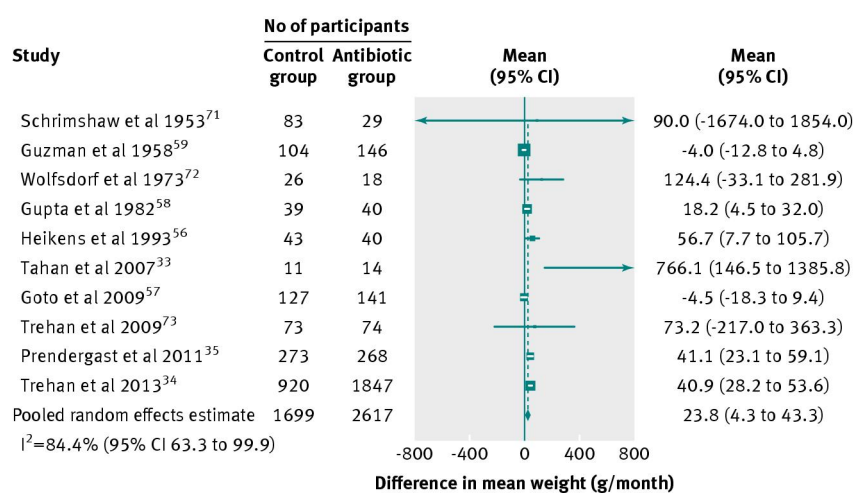


Fig 3 Random effects meta-analyses and forest plots of antibiotic use and weight. Point size reflects study weight