

## S1 Text: Likelihood Derivation.

In this Appendix, we give the log-likelihood for the proposed marginalized two-part model. First, when parameterized in terms of  $p_i$ ,  $\mu_i$  and  $\phi$ , the log-likelihood is

$$l(\mathbf{p}, \boldsymbol{\mu}, \phi | \mathbf{y}) = \sum_{i: y_i > 0} \left\{ \ln(p_i) + g(\mu_i, \phi) + (\mu_i \phi - 1) \ln(y_i) + [(1 - \mu_i)\phi - 1] \ln(1 - y_i) \right\} + \sum_{i: y_i = 0} \ln(1 - p_i), \quad (\text{S1})$$

where

$$g(\mu_i, \phi) = \ln \Gamma(\phi) - \ln \Gamma(\mu_i \phi) - \ln \Gamma((1 - \mu_i)\phi). \quad (\text{S2})$$

Notice the relationship between  $v_i = E(Y_i)$  and  $\mu_i = E(Y_i | Y_i > 0)$ ,

$$v_i = E(Y_i) = p_i E(Y_i | Y_i > 0) = p_i \mu_i. \quad (\text{S3})$$

This gives rise to

$$\mu_i = \frac{v_i}{p_i} = \frac{1 + \exp(-\mathbf{X}_i^T \boldsymbol{\alpha})}{1 + \exp(-\mathbf{X}_i^T \boldsymbol{\gamma})}. \quad (\text{S4})$$

Note also that

$$\begin{aligned} \ln(p_i) &= \mathbf{X}_i^T \boldsymbol{\alpha} - \ln[1 + \exp(\mathbf{X}_i^T \boldsymbol{\alpha})], \\ \ln(1 - p_i) &= -\ln[1 + \exp(\mathbf{X}_i^T \boldsymbol{\alpha})]. \end{aligned} \quad (\text{S5})$$

Thus we can express the log-likelihood in terms of  $\boldsymbol{\alpha}$ ,  $\boldsymbol{\gamma}$  and  $\phi$  :

$$\begin{aligned} l(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi | \mathbf{y}) &= \sum_{i: y_i = 0} -\ln(1 + \exp(\mathbf{X}_i^T \boldsymbol{\alpha})) \\ &+ \sum_{i: y_i > 0} \left\{ \mathbf{X}_i^T \boldsymbol{\alpha} - \ln[1 + \exp(\mathbf{X}_i^T \boldsymbol{\alpha})] + h(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi) \right. \\ &+ \left. \frac{1 + \exp(-\mathbf{X}_i^T \boldsymbol{\alpha})}{1 + \exp(-\mathbf{X}_i^T \boldsymbol{\gamma})} \phi \text{logit}(y_i) + \phi \ln(1 - y_i) - \ln[y_i(1 - y_i)] \right\} \end{aligned} \quad (\text{S6})$$

$$\begin{aligned} &= -\sum_i \ln(1 + \exp(\mathbf{X}_i^T \boldsymbol{\alpha})) \\ &+ \sum_{y_i > 0} \left\{ \mathbf{X}_i^T \boldsymbol{\alpha} + h(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi) + \frac{1 + \exp(-\mathbf{X}_i^T \boldsymbol{\alpha})}{1 + \exp(-\mathbf{X}_i^T \boldsymbol{\gamma})} \phi \text{logit}(y_i) \right. \\ &+ \left. \phi \ln(1 - y_i) - \ln[y_i(1 - y_i)] \right\}, \end{aligned} \quad (\text{S7})$$

where

$$\begin{aligned} h(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi) &\doteq g\{\mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma}), \phi\} \\ &= \ln \Gamma(\phi) - \ln \Gamma[\mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma})\phi] - \ln \Gamma[(1 - \mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma}))\phi]. \end{aligned} \quad (\text{S8})$$

The score equations can be derived straightforwardly:

$$\mathbf{U}_i = \left[ \frac{\partial l_i(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi)}{\partial \boldsymbol{\alpha}}, \frac{\partial l_i(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi)}{\partial \boldsymbol{\gamma}}, \frac{\partial l_i(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi)}{\partial \phi} \right]^T,$$

where

$$\begin{aligned}
\frac{\partial l_i(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi)}{\partial \boldsymbol{\alpha}} &= -\frac{\exp(\mathbf{X}_i^T \boldsymbol{\alpha})}{1 + \exp(\mathbf{X}_i^T \boldsymbol{\alpha})} \mathbf{X}_i^T \\
&+ 1(y_i > 0) \left[ \mathbf{X}_i^T + h'_\alpha(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi) - \frac{\exp(-\mathbf{X}_i^T \boldsymbol{\alpha})}{1 + \exp(-\mathbf{X}_i^T \boldsymbol{\alpha})} \phi \text{logit}(y_i) \mathbf{X}_i^T \right], \\
\frac{\partial l_i(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi)}{\partial \boldsymbol{\gamma}} &= 1(y_i > 0) \left[ h'_\gamma(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi) + \frac{\exp(\mathbf{X}_i^T \boldsymbol{\gamma}) + \exp(\mathbf{X}_i^T \boldsymbol{\gamma} - \mathbf{X}_i^T \boldsymbol{\alpha})}{(1 + \exp(\mathbf{X}_i^T \boldsymbol{\gamma}))^2} \phi \text{logit}(y_i) \mathbf{X}_i^T \right], \\
\frac{\partial l_i(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi)}{\partial \phi} &= 1(y_i > 0) \left[ h'_\phi(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi) + \frac{1 + \exp(-\mathbf{X}_i^T \boldsymbol{\alpha})}{1 + \exp(-\mathbf{X}_i^T \boldsymbol{\gamma})} \text{logit}(y_i) + \ln(1 - y_i) \right],
\end{aligned}$$

and

$$\begin{aligned}
h'_\alpha(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi) &= \frac{\mu'_{i\alpha}(\boldsymbol{\alpha}, \boldsymbol{\gamma})}{\mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma})} - \frac{\mu'_{i\alpha}(\boldsymbol{\alpha}, \boldsymbol{\gamma})}{1 - \mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma})} \\
&+ \sum_{k=1}^{\infty} \left\{ \frac{1}{k + \mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma})\phi} - \frac{1}{k + [1 - \mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma})]\phi} \right\} \phi \mu'_{i\alpha}(\boldsymbol{\alpha}, \boldsymbol{\gamma}), \\
h'_\gamma(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi) &= \frac{\mu'_{i\gamma}(\boldsymbol{\alpha}, \boldsymbol{\gamma})}{\mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma})} - \frac{\mu'_{i\gamma}(\boldsymbol{\alpha}, \boldsymbol{\gamma})}{1 - \mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma})} \\
&+ \sum_{k=1}^{\infty} \left\{ \frac{1}{k + \mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma})\phi} - \frac{1}{k + [1 - \mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma})]\phi} \right\} \phi \mu'_{i\gamma}(\boldsymbol{\alpha}, \boldsymbol{\gamma}), \\
h'_\phi(\boldsymbol{\alpha}, \boldsymbol{\gamma}, \phi) &= \frac{1}{\phi} + \sum_{k=1}^{\infty} \left\{ \frac{\mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma})}{k + \mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma})\phi} + \frac{1 - \mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma})}{k + (1 - \mu_i(\boldsymbol{\alpha}, \boldsymbol{\gamma}))\phi} - \frac{1}{k + \phi} \right\}, \\
\mu'_{i\alpha}(\boldsymbol{\alpha}, \boldsymbol{\gamma}) &= -\frac{\exp(-\mathbf{X}_i^T \boldsymbol{\alpha})}{1 + \exp(-\mathbf{X}_i^T \boldsymbol{\gamma})} \mathbf{X}_i^T, \\
\mu'_{i\gamma}(\boldsymbol{\alpha}, \boldsymbol{\gamma}) &= \frac{\exp(\mathbf{X}_i^T \boldsymbol{\gamma}) + \exp(\mathbf{X}_i^T \boldsymbol{\alpha})}{[1 + \exp(\mathbf{X}_i^T \boldsymbol{\gamma})]^2} \mathbf{X}_i^T.
\end{aligned}$$